The Formation of Inflation Expectations: Micro-data Evidence from Japan*

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Abstract

Using a large panel dataset in Japan, we examine the formation of inflation expectations by households. It is widely known that the expectation formation of inflation rates are not necessarily "rational". Rather, forecasts of inflation rates by households are biased upward. The literature shows that the upward bias is explained by socio-economic factors as well as information rigidities. This study investigates whether not only respondent's attributes but also the hypothesis of sticky information determine the bias of households' inflation forecasts. In order to uncover the determinants of the forecast bias, we conduct online survey about inflation outlook to about 50,000 households to collect the data about the inflation outlook and the frequency of updating information about inflation rates. There are three findings. First, the revision of inflation expectations over both shorter and longer horizons respond to changes in food prices not those of oil prices as in Coibion and Gorodnichenko (2015b). It is suggested that perceived inflation rates in food prices may affects inflation expectations of households. Second, about half of households update their inflation set every quarter. The fraction of households who revise their information about inflation rates is larger than that reported in Carroll (2003). Third, the upward bias of inflation expectations shrinks when respondents update their information set.

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information rigidities; sticky information

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1 Introduction

It is widely known that the expectation formation of inflation rates are not necessarily "rational". Rather, forecasts of inflation rates by households are biased upward. The literature shows that the upward bias is explained by socio-economic factors as well as information rigidities. This study investigates whether not only respondent's attributes but also the hypothesis of sticky information determine the bias of households' inflation forecasts.

2 Data

We conduct online survey about inflation outlook to about 50,000 households and ask how often they update the frequency of updating information about inflation rates. Respondents are asked to answer the price levels after one, three, and ten years if the price level is 10,000 today. The data covers from 2015:Q4 to 2018:Q2. Tables 1 and 2 show the basic statistics of inflation expectations and the fraction of households who update their information about inflation rates. Table 2 shows that about half of households revise their information set every quarter and the fraction is larger than that reported in Carroll (2003). The fact that not all respondents do not update their information set about inflation rates supports the sticky information hypothesis.

3 Estimation Strategy and Results

In order to examine the determinants of the upward bias of inflation expectations, we first regress the gap between inflation expectations between households and professionals on changes in food and oil prices. Following Coibion and Gorodnichenko (2015b), we estimate the following equation:

$$E_t[\pi_{t,t+4}^{Households}] - E_t[\pi_{t,t+4}^{Professionals}] = \beta \times \pi_{t-k,t-1} + \varepsilon_t,$$

where $E_t[\pi_{t,t+4}]$ and $\pi_{t-k,t-1}$ are denoted as inflation forecasts over the four-quarter-ahead and changes of energy prices or food prices from t - k to t - 1. Table 2 shows that the gap between inflation expectations of households and professionals is determined by not energy price changes but food price changes.

This result is supported by panel data. We regress respondent *i*'s revisions $(E_t^i[\pi_{t,t+4}] - E_{t-2}^i[\pi_{t-2,t+2}])$ on $\pi_{t-2,t}^{Energy}$ and $\pi_{t-2,t}^{Food}$. Table 4 reports that households' forecasts respond more strongly to food price changes than energy price changes, significantly.

We also test the sticky information hypothesis. If the sticky information hypothesis holds, the bias of respondents who update their information set should shrink. We regress the individual forecasters' bias from median on socio-economic factors and the intersection of the factors with the dummy variable (D^{Update}) which takes one if information is updated, otherwise zero. Table 5 shows that while socio-economic factors such as sex, age, income, education, and marital status explain the upward bias of inflation expectations, the sings of the intersections with D^{Update} are negative and significant in almost all cases. This result implies that the bias shrinks when information are revised. These findings are consistent with Ehrmann et al. (2017) and support the sticky information hypothesis.

References

- Carroll, C.D. (2003) "Macroeconomic Expectations of Households and Professional Forecasters", *Quarterly Journal of Economics*, Vol. 118, No. 1, pp.269–298.
- Coibion, O. and Y. Gorodnichenko (2012) "What Can Survey Forecasts Tell Us about Information Rigidities?", *Journal of Political Economy*, Vol. 120, No. 1, pp.116–159.
- Coibion, O. and Y. Gorodnichenko (2015) "Information Rigidity and the Expectations Formation Process: A Simple Framework and New Facts", *American Economic Review*, Vol. 105, No. 8, pp.2644–2678.
- Coibion, O. and Y. Gorodnichenko (2015) "Is the Phillips Curve Alive and Well after All? Inflation Expectations and the Missing Disinflation", *American Economic Journal: Macroeconomics*, Vol. 7, No. 1, pp.197–232.
- Ehrmann, M., D. Pfajfar, and E. Santoro (2017) "Consumers' Attitudes and Their Inflation Expectations", *International Journal of Central Banking*, Vol. 13, No. 1, pp.225–259.
- Mankiw, N. G. and R. Reis (2002) "Sticky Information Versus Sticky Prices: A Proposal to Replace the New Keynesian Phillips Curve", *Quarterly Journal of Economics*, Vol. 117, No. 4, pp.1295–1328.
- Newey, W. K. and K. D. West (1987) "A Simple Positive Semi-Definite, Heteroskedasticity and Autocorrelation Consistent Covariance Matrix", *Econometrica*, Vol. 55, No. 3, pp.703–708.

	1-year average			3-year average			10-year average		
	Mean	Median	Obs.	Mean	Median	Obs.	Mean	Median	Obs.
All	2.5%	0.5%	106,079	2.1%	0.9%	107,008	1.4%	0.9%	106,783
Woman	2.8%	0.9%	51,209	2.3%	0.9%	51,439	1.6%	0.9%	51,224
Man	2.3%	0.5%	54,870	1.8%	0.6%	55,569	1.3%	0.9%	55,559
L. Edu.	2.8%	0.9%	47,654	2.3%	0.9%	47,990	1.6%	0.9%	47,863
H. Edu.	2.3%	0.5%	58,425	1.9%	0.6%	59,018	1.3%	0.9%	58,920
L. Income	2.7%	0.5%	63,758	2.2%	0.9%	64,286	1.5%	0.9%	64,145
H. Income	2.3%	0.5%	42,321	1.9%	0.6%	42,722	1.3%	0.9%	42,638
Not Married	2.8%	0.5%	28,987	2.3%	0.9%	29,385	1.6%	0.9%	29,584
Married	2.4%	0.5%	77,092	2.0%	0.8%	77,623	1.4%	0.9%	77,199
Updated	2.6%	0.5%	86,022	2.1%	0.9%	86,737	1.4%	0.9%	86,358
Not Updated	2.4%	0.2%	20,057	2.1%	0.6%	20,271	1.5%	0.9%	20,425

Table 1: The basic statistics of inflation expectations

Note: "Lower", "Higher", and "Education" are abbreviated as "L", "H", and "Edu", respectively.

Table 2: The fraction of households who update information about inflation rates

	Male	F 1	Total
	widte	Female	1000
Not Undeted	59,035	94,333	153,368
Not Opualed	44.23%	53.08%	49.29%
Undeted	74,426	83,382	157,808
Opualeu	55.77%	46.92%	50.71%
Total	133,461	177715	311,176
10141	100.00%	100.00%	100.00%
	Not Updated Updated Total	Not Updated 44.23% Updated 74,426 55.77% 133,461	Not Updated 44.23% 53.08% Updated 74,426 83,382 55.77% 46.92% Total 133,461 177715

	$E_t[\pi]$	[Households]	$-E_t[\pi_{t,t+4}^{Prof}]$	$E_t[\pi_{t,t+4}^{Professionals}] = \beta \times \pi_{t-k,t-1}$				
	(1)	(2)	(3)	(4)	(5)	(6)		
$\pi^{Energy}_{t-2,t-1}$	0.0530				0.0287			
,	(0.0958)				(0.0847)			
$\pi^{Energy}_{t-3,t-1}$		0.0499				0.0298		
		(0.0574)				(0.0498)		
$\pi^{Food}_{t-2,t-1}$			0.572***		0.559***			
- J			(0.185)		(0.195)			
$\pi^{Food}_{t-3,t-1}$				0.466***		0.449***		
				(0.127)		(0.134)		
Observations	54	53	54	53	54	53		

Table 3: Determinants of the bias between households and professionals

Note: *** indicates 1% significance. Standard errors in parentheses are calculated by the Newey-West (1987) estimator. The data covers from 2004Q2 to 2017Q3.

	$E_t^i[\pi_{t,t+4}]$ -	$-E_{t-2}^{i}[\pi_{t-2,t+}]$	$\alpha_2] = \alpha + \beta_1 \times \beta_1 \times \beta_2$	$\pi_{t-2,t}^{Energy} + \beta_2 >$	$\times \pi^{Food}_{t-2,t} + \varepsilon^i_t$
		"Spot"		"Forv	ward"
	1 year	3 year	10 year	1–3 year	3–10 year
α	-0.745^{***}	-0.516^{***}	-0.270 ***	-0.425^{***}	-0.168***
	(0.033)	(0.022)	(0.015)	(0.022)	(0.013)
$\beta_1: \pi_{t-2,t}^{Energy}$	0.121***	0.069***	0.030***	0.044***	0.017***
	(0.009)	(0.006)	(0.004)	(0.006)	(0.004)
β_2 : $\pi^{Food}_{t-2,t}$	0.183***	0.141***	0.066***	0.125***	0.037***
,	(0.016)	(0.011)	(0.008)	(0.012)	(0.007)
F-test	13.597	38.326	16.592	44.511	7.490
$(H_0:\beta_1=\beta_2)$	10.077	20.220	10.072	11011	71120
p-value	0.000	0.000	0.000	0.000	0.006
Observation	37,537	38,103	37,745	37,111	36,856

Table 4: Do revisions respond to price changes? Joint Test

Note: Standard errors in parentheses are clustered at individual levels, and *** indicates 1% significance.

	1-year	3-year	10-year	1 to 3-year	3 to 10-yea
$\beta: \pi^{Food}_{t-2,t}$	0.129***	0.0530***	0.0151***	0.0765***	0.00964*
v = 2, v	(0.00878)	(0.00659)	(0.00499)	(0.00662)	(0.00382
Woman	0.745***	0.583***	0.371***	0.489***	0.282**
	(0.0763)	(0.0592)	(0.0451)	(0.0561)	(0.032
Age (< 50 years old)	0.359***	0.196***	0.0783*	0.234***	0.0621*
	(0.0695)	(0.0529)	(0.0409)	(0.0517)	(0.028)
Lower Income	0.539***	0.322***	0.243***	0.343***	0.193**
	(0.0701)	(0.0530)	(0.0404)	(0.0506)	(0.0290
Lower Education	0.680***	0.446***	0.239***	0.385***	0.184**
	(0.0755)	(0.0598)	(0.0443)	(0.0567)	(0.032
Not Married	0.669***	0.548***	0.323***	0.506***	0.203**
	(0.0864)	(0.0701)	(0.0512)	(0.0672)	(0.036
D^{Update}	1.170***	0.485***	0.180***	0.546***	0.287**
	(0.0413)	(0.0308)	(0.0233)	(0.0291)	(0.018)
$D^{Update} \times \operatorname{Woman}$	-0.223^{***}	-0.155 **	-0.123 **	-0.207^{***}	-0.174**
	(0.0833)	(0.0635)	(0.0477)	(0.0606)	(0.034)
$D^{Update} \times Age$	-0.330***	-0.126**	0.00293	-0.193***	-0.0625
	(0.0773)	(0.0580)	(0.0441)	(0.0566)	(0.031)
$D^{Update} \times$ Income	-0.223^{***}	-0.0867	-0.0843*	-0.159***	-0.105**
	(0.0774)	(0.0579)	(0.0436)	(0.0554)	(0.032
$D^{Update} \times \text{Education}$	-0.349***	-0.227***	-0.0789*	-0.233^{***}	-0.0914**
	(0.0827)	(0.0642)	(0.0473)	(0.0613)	(0.035)
$D^{Update} \times \text{Not Married}$	-0.223 **	-0.182^{**}	-0.0467	-0.213^{***}	-0.0875*
	(0.0951)	(0.0756)	(0.0555)	(0.0726)	(0.040
Observations	106,079	107,008	106,783	104,857	104,64

Table 5: The determinants of bias from median $(\bar{\pi}_{i,t}^e)$